

On the Identifiability of a Two-Factor Model with Correlated Residuals

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Consider the following two-factor model:

$$\begin{aligned} X_j^1 &= \theta_j^1 H_1 + Y_j^1, & j = 2, \dots, n_1, \\ X_j^2 &= \theta_j^2 H_2 + Y_j^2, & j = 2, \dots, n_2, \end{aligned} \quad (1)$$

where $X = (X_1^1, \dots, X_{n_1}^1, X_1^2, \dots, X_{n_2}^2)$ is a vector of observable variables with a vector of means m_X and covariance matrix Σ_X ; $\theta_X = (\theta_1^1, \dots, \theta_{n_1}^1, \theta_1^2, \dots, \theta_{n_2}^2)$ is a vector of parameters; $H = (H_1, H_2)$ is a vector of latent variables normally distributed with $E(H_1) = E(H_2) = 0$ and $Var(H_1) = Var(H_2) = 1$; $Y = (Y_1^1, \dots, Y_{n_1}^1, Y_1^2, \dots, Y_{n_2}^2)$ is a vector of residuals normally distributed with positive definite covariance matrix; Y and H are independent.

We will assume that probabilistic relationships among components of the vector Y are represented by a Bayesian network [1] with a structure S_Y , that is a directed acyclic graph with nodes identified with components of the vector Y , and a vector of parameters θ_Y . The model (1) is a generalization of the single-factor model with correlated residuals suggested in [2].

The model (1) is said to be identifiable if the vector of parameters $\theta = (\theta_X, \theta_Y)$ can be uniquely determined by m_X and Σ_X .

Transform the structure S_Y by the following way. Connect by edges all disconnected by arcs nodes and then delete all arcs. Obtained graph \bar{S}_Y is called complementary graph of the structure S_Y .

Let C_1 be a cycle that consists of nodes Y_1, \dots, Y_k and C_2 be a graph that consists of a cycle with nodes Y_1, \dots, Y_{i_1} , a simple chain with nodes Y_{i_1}, \dots, Y_{i_2} and a cycle with nodes Y_{i_2}, \dots, Y_n .

Define two functions:

$$F_1(C_1) = \sum_{i=1}^{k-1} \chi(Y_i, Y_{i+1})(-1)^{i+1} + \chi(Y_1, Y_k)(-1)^{k+1}$$

and

$$\begin{aligned} F_2(C_2) &= \sum_{j=1}^{i_1-1} \chi(Y_j, Y_{j+1})(-1)^{j+1} + \chi(Y_1, Y_{i_1})(-1)^{i_1+1} + \\ &\quad 2\left(\sum_{j=i_1}^{i_2-1} \chi(Y_j, Y_{j+1})(-1)^{j+1}\right) + \\ &\quad \sum_{j=i_2}^{n-1} \chi(Y_j, Y_{j+1})(-1)^{j+1} + \chi(Y_n, Y_{i_2})(-1)^{i_2}, \end{aligned}$$

where $\chi(Y_i, Y_j) = 0$, if $Y_i, Y_j \in \{Y_1^1, \dots, Y_{n_1}^1\}$ or $Y_i, Y_j \in \{Y_1^2, \dots, Y_{n_2}^2\}$ and $\chi(Y_i, Y_j) = 1$ in opposite case.

The following theorem provides identifiability conditions for model (1).

Theorem. *Assume that Σ_X^{-1} does not contain zero elements and $\theta_1^1 > 0, \theta_1^2 > 0$. A necessary and sufficient condition for model (1) to be identified is that every connectivity component of the graph \bar{S}_Y contains at least one odd cycle and at least one connectivity component contains either even cycle C_1 for that $F_1(C_1) \neq 0$ or subgraph C_2 that consists of two odd cycles connected by simple chain for that $F_2(C_2) \neq 0$.*

REFERENCES

1. Heckerman D. A tutorial on learning Bayesian networks. Tech. Report MSR-TR-95-06, Microsoft, WA, 1995.
2. Vicard P. On the identification of a single-factor model with correlated residuals // *Biometrika*, 2000, v. 87, p. 199-205.